

ISyE 3770 Assignment 5: Point Estimator

Due date: 11:59 PM, Friday, March 29, 2024.

Question 1 (MSE). Let X_1, X_2 be independent random variables with mean μ and variance σ^2 . Suppose we have two estimators of μ :

$$\hat{\Theta}_1 = \frac{X_1 + X_2}{2}$$

$$\hat{\Theta}_2 = \frac{X_1 + 3X_2}{4}.$$

- 1) Are both estimators unbiased? (8 points)
- 2) What is the variance of each estimator? (8 points)
- 3) What is the MSE of two estimators? (8 points)

Solution. 1) Yes. It can be verified that $\mathbb{E}[\hat{\Theta}_1] = \mu$ and $\mathbb{E}[\hat{\Theta}_2] = \mu$.

2) You can check that

$$\text{Var}(\hat{\Theta}_1) = \frac{1}{4}\text{Var}(X_1) + \frac{1}{4}\text{Var}(X_2) = \frac{1}{2}\sigma^2$$

and

$$\text{Var}(\hat{\Theta}_2) = \frac{1}{16}\text{Var}(X_1) + \frac{9}{16}\text{Var}(X_2) = \frac{5}{8}\sigma^2.$$

3) Recall that $\text{MSE} = (\text{Bias})^2 + \text{Variance}$. Since these two estimators are unbiased, it holds that

$$\text{MSE}(\hat{\Theta}_1) = \text{Var}(\hat{\Theta}_1) = \frac{\sigma^2}{2}, \quad \text{MSE}(\hat{\Theta}_2) = \text{Var}(\hat{\Theta}_2) = \frac{5\sigma^2}{8}.$$

□

Question 2 (MSE). Suppose $X \sim \text{Uni}(\theta, 3\theta)$ with $\theta > 0$. Let X_1, \dots, X_n be n i.i.d. random variables following the same distribution as X .

- 1) Prove that $\frac{\bar{X}}{2}$ is an unbiased estimator of θ . (8 points)
- 2) Calculate the MSE of $\frac{\bar{X}}{2}$ and \bar{X} . (8 points)

Solution. 1) You can check that $\mathbb{E}[\frac{\bar{X}}{2}] = \frac{1}{2}\mathbb{E}[X] = \theta$.

2) Recall that $\text{MSE} = (\text{Bias})^2 + \text{Variance}$. Therefore,

$$\text{MSE}(\frac{\bar{X}}{2}) = \text{Var}(\frac{\bar{X}}{2}) = \frac{1}{4} \cdot \frac{\text{Var}(X)}{n} = \frac{1}{4} \frac{4\theta^2}{12n} = \frac{\theta^2}{12n}.$$

Moreover, $\text{Bias}(\bar{X}) = \theta$ and therefore,

$$\text{MSE}(\bar{X}) = \theta^2 + \text{Var}(\bar{X}) = \theta^2 + \frac{\theta^2}{3n} = (1 + \frac{1}{3n})\theta^2.$$

□

Question 3 (Maximum Likelihood Estimator). A random variable X has the following probability density function:

$$f(x; \theta) = \frac{1}{2\theta^3} x^2 e^{-x/\theta}, \quad 0 < x < \infty, 0 < \theta < \infty.$$

Let X_1, \dots, X_n be n i.i.d. random variables following the same distribution as X . Find the maximum likelihood estimator for θ . (15 points)

Solution. • Step 1: write down the log-likelihood function and simplify it:

$$\begin{aligned} l(\theta; X_1, \dots, X_n) &= \sum_{i=1}^n \log f(X_i; \theta) = \sum_{i=1}^n \log \left[\frac{1}{2\theta^3} X_i^2 e^{-X_i/\theta} \right] \\ &= \sum_{i=1}^n \left[-3 \log \theta - \frac{X_i}{\theta} \right] + \text{Constant} \\ &= -3n \log \theta - \frac{\sum_{i=1}^n X_i}{\theta} + \text{Constant}. \end{aligned}$$

• Step 2: find the maximizer of $l(\theta; X_1, \dots, X_n)$: let $\hat{\theta}$ be the maximizer of $l(\theta; X_1, \dots, X_n)$, then

$$\frac{\partial l(\hat{\theta}; X_1, \dots, X_n)}{\partial \theta} = \frac{-3n}{\hat{\theta}} + \frac{\sum_{i=1}^n X_i}{\hat{\theta}^2} = 0 \implies \hat{\theta} = \frac{\bar{X}}{3}.$$

□

Question 4 (Maximum Likelihood Estimator). A random variable X has the following probability density function:

$$f(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right).$$

Let X_1, \dots, X_n be n i.i.d. random variables following the same distribution as X . Find the maximum likelihood estimator for μ and σ^2 . (15 points)

Solution. • Step 1: write down the log-likelihood function and simplify it:

$$\begin{aligned} l(\mu, \sigma^2; X_1, \dots, X_n) &= \sum_{i=1}^n \log f(X_i; \mu, \sigma^2) = \sum_{i=1}^n \log \left[\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \right] \\ &= \sum_{i=1}^n \left[-\frac{1}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} (X_i - \mu)^2 \right] + \text{Constant} \\ &= -\frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 + \text{Constant}. \end{aligned}$$

• Step 2: find the maximizer of $l(\mu, \sigma^2; X_1, \dots, X_n)$: let $\hat{\mu}, \hat{\sigma}^2$ be the maximizer of $l(\mu, \sigma^2; X_1, \dots, X_n)$, then

$$\frac{\partial l(\hat{\mu}, \hat{\sigma}^2; X_1, \dots, X_n)}{\partial \mu} = \frac{1}{\hat{\sigma}^2} \sum_{i=1}^n (X_i - \hat{\mu}) = 0 \implies \hat{\mu} = \bar{X},$$

and

$$\frac{\partial l(\hat{\mu}, \hat{\sigma}^2; X_1, \dots, X_n)}{\partial \sigma^2} = -\frac{n}{2\hat{\sigma}^2} + \frac{1}{2\hat{\sigma}^4} \sum_{i=1}^n (X_i - \hat{\mu})^2 = 0 \implies \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2.$$

□

Question 5 (Maximum Likelihood Estimator and Method of Moment Estimator). Let X be an exponential random variable with parameter λ . Let X_1, \dots, X_n be n i.i.d. random variables following the same distribution as X .

- 1) Find the maximum likelihood estimator for λ . (15 points)
- 2) Find the method of moment estimator for λ . (15 points)

Solution. 1) • Step 1: write down the log-likelihood function and simplify it:

$$\begin{aligned} l(\lambda; X_1, \dots, X_n) &= \sum_{i=1}^n \log f(X_i; \lambda) = \sum_{i=1}^n \log [\lambda e^{-\lambda X_i}] \\ &= \sum_{i=1}^n [\log \lambda - \lambda X_i] + \text{Constant} \\ &= n \log \lambda - \lambda \sum_{i=1}^n X_i + \text{Constant}. \end{aligned}$$

- Step 2: find the maximizer of $l(\lambda; X_1, \dots, X_n)$: let $\hat{\lambda}$ be the maximizer of $l(\lambda; X_1, \dots, X_n)$, then

$$\frac{\partial l(\hat{\lambda}; X_1, \dots, X_n)}{\partial \lambda} = \frac{n}{\hat{\lambda}} - \sum_{i=1}^n X_i = 0 \implies \hat{\lambda} = \frac{1}{\bar{X}}.$$

- 2) Based on the method of moment, we find the sample mean equals \bar{X} , and the population mean equals $\frac{1}{\lambda}$.

Therefore,

$$\bar{X} = \frac{1}{\lambda} \implies \tilde{\lambda} = \frac{1}{\bar{X}}.$$

□