

Chapter 4 **Bivariate** Distributions (二元分布)

Section 4.1 **Bivariate** Distributions with the **discrete** type

➤ Motivation

very often, the outcome of a random experiment is a **tuple** of several things of interests:

- Observe female college students to obtain information such as *height* x , and *weight* y .
- Observe high school students to obtain information such as *rank* x , and *score of college entrance examination* y .

➤ In order to define joint probability mass function (joint pmf):

* Complete way :

① identify the Sample Space S ;

② Define a *RV* $Z = \begin{bmatrix} X \\ Y \end{bmatrix} : S \rightarrow Z(S)$;

③ Define a *pmf* for Z , $f(z) : Z(S) \rightarrow [0,1]$.

* Simplified way :

① Ignore the Sample Space S ;

② Specify $Z(S)$ directly and denote it by D ;

③ Define the *pmf* for Z , $f(z) : D \rightarrow [0,1]$;

equivalently, for $\begin{bmatrix} X \\ Y \end{bmatrix}$, $f(x, y) : D \rightarrow [0,1]$.

Definition 4.1-1 [joint probability mass function (joint pmf)]

Let X and Y be 2 RV s. The probability that $X = x$ and $Y = y$ is denoted by $f(x, y) = P(X = x, Y = y)$.

The function $f(x, y): D \rightarrow [0,1]$ is called the **joint probability mass function (joint pmf)** of X if:

$$\textcircled{1} 0 \leq f(x, y) \leq 1;$$

$$\textcircled{2} \sum_{(x,y) \in D} f(x, y) = 1;$$

$$\textcircled{3} P[(X, Y) \in A] \triangleq P(\{(x, y) \in A\}) = \sum_{(x,y) \in A} f(x, y), \quad A \subseteq D.$$

Example 1 [Page 134]

Roll a pair of fair dice. The sample Space contains 36 outcomes. And let X denote the smaller outcome and Y the larger outcome on the die.

For instance, if the outcome is (3,2), then $X=2, Y=3$.

Obviously, $P(\{X = 2, Y = 3\}) = 1/36 + 1/36 = 2/36$.

$P(\{X = 2, Y = 2\}) = 1/36$.

Furthermore, the *joint pmf* of X and Y is: $f(x, y) = \begin{cases} 1/36, & 1 \leq x = y \leq 6 \\ 2/36, & 1 \leq x < y \leq 6 \end{cases}$

Definition 4.1-2 [Marginal pmf]

Let X and Y have the joint probability mass function $f(x, y) : D \rightarrow [0,1]$. Sometimes we are interested in the pmf of X or Y alone, which is called the **marginal probability mass function of X or Y** and defined by

$$f_X(x) = \sum_{y \in D_Y} f(x, y) = P(X = x), \quad x \in D_X = \{\text{all possible values of } X \text{ in } D\}.$$

$$f_Y(y) = \sum_{x \in D_X} f(x, y) = P(Y = y), \quad y \in D_Y = \{\text{all possible values of } Y \text{ in } D\}.$$

Definition 4.1-3 [independent Random Variables]

The random variables X and Y are **independent** if and only if, for every $x \in D_X$ and $y \in D_Y$,

$$P(\underline{X} = \underline{x}, \underline{Y} = \underline{y}) = P(\underline{X} = \underline{x})P(\underline{Y} = \underline{y})$$

or equivalently,

$A \cap B$

Event A

Event B

$$f(x, y) = f_X(x)f_Y(y).$$

otherwise, X and Y are said to be **dependent**.

Example 2 [Page 135]

Let the joint *pmf* of X and Y be defined by

$$f(x, y) = \frac{x + y}{21}, \quad x = 1, 2, 3, \quad y = 1, 2.$$

Check if *RV* X and Y are independent.

Solution :

$$f_X(x) = \sum_{y \in D_Y} f(x, y) = \sum_{y=1}^2 \frac{x + y}{21} = \frac{2x + 3}{21}, \quad x = 1, 2, 3.$$

$$f_Y(y) = \sum_{x \in D_X} f(x, y) = \sum_{x=1}^3 \frac{x + y}{21} = \frac{3y + 6}{21}, \quad y = 1, 2.$$

$$f(x, y) = \frac{x + y}{21} \neq \frac{2x + 3}{21} \cdot \frac{3y + 6}{21} = f_X(x)f_Y(y) \Rightarrow X \text{ and } Y \text{ are dependent.}$$

What's the interpretation of $f_X(x)$ and $f_Y(x)$ and independence?

Consider the *conditional pmf* :

$$f(y|x) = P(Y = y|X = x) = \frac{f(x, y)}{f_X(x)}; f(x|y) = P(X = x|Y = y) = \frac{f(x, y)}{f_Y(x)}$$

*We will
learn it
formally in
section 4.3*

➤ Expectation and Value

Let X_1 and X_2 be discrete *RV* with their joint *pmf* $f(x_1, x_2) : D \rightarrow [0,1]$. Consider a function $u(x_1, x_2)$ of x_1 and x_2 . Then:

Expectations of functions of bivariate RVs are computed just as with univariate RVs.

(a) The **mathematical expectation** of $u(X_1, X_2)$, if exists, is given by

$$E[u(X_1, X_2)] = \sum_{(x_1, x_2) \in D} u(x_1, x_2) f(x_1, x_2).$$

(b) If $u_i(X_1, X_2) = X_i$ for $i = 1, 2$, then

$$E(X) = \sum_{(x,y) \in \bar{S}} xf(x,y) = \sum_{x \in \bar{S}_X} xf_X(x).$$

$$E[u_i(X_1, X_2)] = E(X_i) = u_i$$

is called the **mean** of X_i for $i = 1, 2$.

(c) If $u_i(X_1, X_2) = (X_i - u_i)^2$ for $i = 1, 2$, then

$$E[u_i(X_1, X_2)] = E[(X_i - u_i)^2] = \sigma_i^2 = \text{Var}(X_i)$$

is called the **variance** of X_i for $i = 1, 2$.

Example 1 [Page 134]—revisited

Recall that X and Y are discrete *RVs* with joint *pmf*
 $f(X, Y) : D \rightarrow [0, 1]$ with $D_X = D_Y = \{1, 2, 3, 4, 5, 6\}$

$$f(x, y) = \begin{cases} 2/36, & 1 \leq x < y \leq 6 \\ 1/36, & 1 \leq x = y \leq 6 \end{cases}$$

Compute $E(X+Y)$:

Solution :

$$\begin{aligned} E(X+Y) &= \sum_{(x,y) \in D} (x+y) f(x,y) = \sum_{1 \leq x=y \leq 6} (x+y) \cdot \frac{1}{36} + \sum_{1 \leq x < y \leq 6} (x+y) \frac{2}{36} \\ &= \sum_{x=1}^6 2x \cdot \frac{1}{36} + \sum_{x=1}^6 \sum_{y=x+1}^6 (x+y) \cdot \frac{2}{36} = \frac{252}{36}. \end{aligned}$$



Work it by yourself!

Chapter 4

Bivariate Distributions (二元分布)

Section 4.2

The correlation coefficient

Let X_1 and X_2 be discrete RV with their joint pmf $f(x_1, x_2) : D \rightarrow [0,1]$.

recall that for $u(X, Y)$, its expectation $E[u(X, Y)] = \sum_{(x,y) \in D} u(x, y) f(x, y)$.

Definition 4.2-1 [Covariance of X and Y]

Take $u(X, Y) = [X - E(X)][Y - E(Y)]$

$$E[(X - E(X))(Y - E(Y))] = Cov(X, Y),$$

which is called the covariance of X and Y .

➤ Motivation: To study the relation between X and Y .

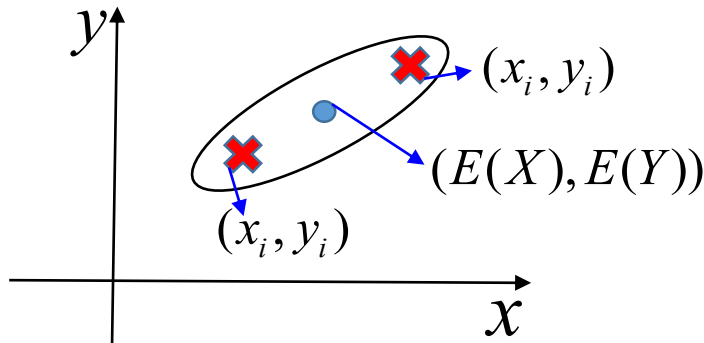
- $Cov(X, Y) = E(XY) - E(X)E(Y)$ → Verify it by yourself!

When $Cov(X, Y) = 0$, we say X and Y are **uncorrelated**.

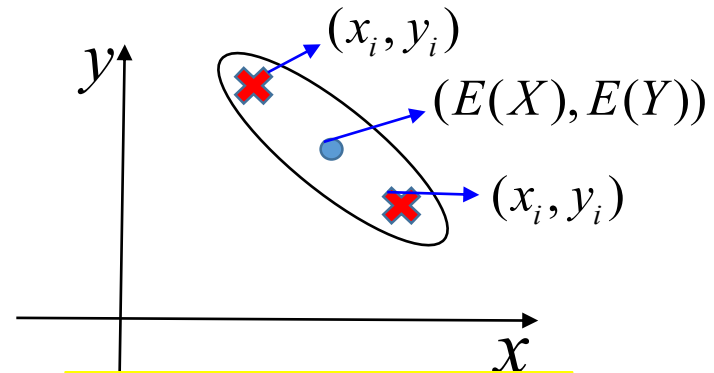
- Interpretation: Roughly speaking, a positive or negative covariance indicates that the values of $X - E(X)$ and $Y - E(Y)$ obtained in a single experiment 'tend' to have the **same** or the **opposite** sign.

Example 1: Demonstration of positively correlated and negatively correlated RVs

Assume that X and Y are uniformly distributed over the ellipses.



positively correlated



negatively correlated

Independence of X and Y could imply the **uncorrelation** of X and Y .

Consider the case that X and Y are independent:

$$\begin{aligned}
 E(XY) &= \sum_{(x,y) \in D} xyf(x,y) = \sum_{x \in D_X} \sum_{y \in D_Y} xyf_X(x)f_Y(y) \\
 &= \sum_{x \in D_X} xf_X(x) \left[\sum_{y \in D_Y} yf_Y(y) \right] = E(X)E(Y).
 \end{aligned}$$

$$f(x,y) = f_X(x)f_Y(y)$$

$$\Rightarrow D = D_X D_Y$$

Therefore, $cov(X,Y) = E(XY) - E(X)E(Y) = 0$.

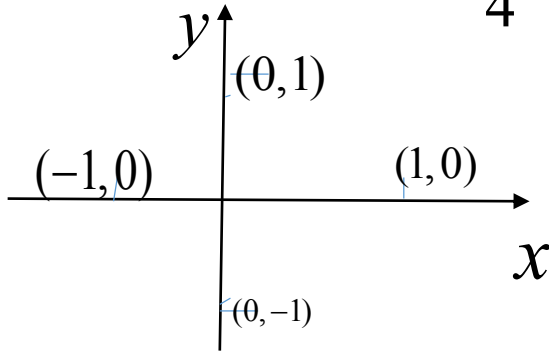
Independent of 2 RVs \Rightarrow uncorrelation of 2 RVs.

However, the converse is not true, that is to say, there exists X and Y which are **uncorrelated** but **not independent**.

Example 2 (**uncorrelation doesn't imply independence**)

Let X and Y be RVs that take values $(1, 0)$, $(0, 1)$, $(-1, 0)$, $(0, -1)$

and with probability $\frac{1}{4}$, as shown in the figure below.



Q1 : what are the **marginal pmf** of X and Y ?

Q2 : what is $Cov(X, Y)$?

Q3 : Are X and Y *independent*?

Solution : To find marginal pmf of X and Y , $D_X = D_Y = \{-1, 0, 1\}$.

$$f_X(x) = \begin{cases} 1/4, & x = 1 \\ 1/2, & x = 0 \\ 1/4, & x = -1 \end{cases}, \quad f_Y(y) = \begin{cases} 1/4, & y = 1 \\ 1/2, & y = 0 \\ 1/4, & y = -1 \end{cases}$$

$$Cov(X, Y) = E(XY) - E(X)E(Y) = 0 - 0 \cdot 0 = 0.$$

$$f_X(0)f_Y(1) = \frac{1}{2} \cdot \frac{1}{4} = \frac{1}{8} \neq f(0, 1) = \frac{1}{4} \Rightarrow X \text{ and } Y \text{ are not independent!}$$

Definition 4.2-2 [correlation coefficients]

The correlation coefficients of X and Y that have nonzero variance is defined as

$$\rho(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}.$$

- It is a normalized version of $Cov(X, Y)$ and in fact $-1 \leq \rho \leq 1$
- Interpretation: $\rho > 0$ (or $\rho < 0$) indicate the values of $X - E(X)$ and $Y - E(Y)$ 'tend' to have the same (or opposite, respectively) sign.
- $\rho > 0$ (or $\rho < 0$) have the same interpretation as $Cov(X, Y) > 0$ (or $Cov(X, Y) < 0$)
- The size of $|\rho|$ provides a normalized measure of the extent to which this is true.
- $\rho = 1$ or $\rho = -1$ if and only if there exists a positive (or negative, respectively) constant c such that

$$Y - E(Y) = c[X - E(X)]$$

Example 3

Consider n **independent** tosses of a coin with probability of a head equal to p . Let X and Y be the number of heads and of tails, respectively. Calculate the correlation coefficient of X and Y .

Solution:

$$X + Y = n \Rightarrow E(X) + E(Y) = n \Rightarrow X - E(X) = -[Y - E(Y)]$$

$$\text{Cov}(X, Y) = E[(X - E(X))(Y - E(Y))] = -E[(Y - E(Y))^2] = -\text{Var}(Y)$$

$$\text{Var}(X) = E[(X - E(X))^2] = E[(Y - E(Y))^2] = \text{Var}(Y)$$

$$\Rightarrow \rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}} = \frac{-\text{Var}(Y)}{\sqrt{\text{Var}(Y)}\sqrt{\text{Var}(Y)}} = -1.$$